

Solution 1.4: we have by definition

$$\mathbb{E}(T_n - \mu)^2 = \mathbb{E} \left[ (\bar{X}_n - \mu)^2 1_{|\bar{X}_n| > n^{-1/4}} \right] + \mu^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right).$$

At  $\mu = 0$ , it simplifies to

$$\mathbb{E}T_n^2 = \mathbb{E} \left[ \bar{X}_n^2 1_{|\bar{X}_n| > n^{-1/4}} \right].$$

The boundedness of  $\bar{X}_n$  and Cauchy-Schwarz's inequality lead to

$$\mathbb{E}T_n^2 = \mathbb{E} \left[ \bar{X}_n^2 1_{|\bar{X}_n| > n^{-1/4}} \right] \leq \sqrt{\mathbb{E} \left[ \bar{X}_n^4 \right] \mathbb{P} \left( |\bar{X}_n| > n^{-1/4} \right)} \leq 2b^2 \exp \left( -\frac{\sqrt{n}}{4b^2} \right). \quad \square$$

5. Show that for any fixed  $\mu \neq 0$ ,  $R(T_n, \mu)$  converges to zero at rate  $O(1/n)$ .

Solution: thanks to the triangle inequality  $|\bar{X}_n - \mu| \geq |\mu| - |\bar{X}_n|$ , we deduce

$$\mathbb{P} \left( |\bar{X}_n| \leq |\mu| - n^{-1/4} \right) \leq \mathbb{P} \left( |\bar{X}_n - \mu| \geq n^{-1/4} \right) \leq 2 \exp \left( -\frac{\sqrt{n}}{2b^2} \right).$$

When  $\mu \neq 0$ , as soon as  $|\mu| > 2n^{-1/4}$ , we obtain

$$\mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right) \leq \mathbb{P} \left( |\bar{X}_n| \leq |\mu| - n^{-1/4} \right) \leq 2 \exp \left( -\frac{\sqrt{n}}{2b^2} \right).$$

Again, writing out the expression for the quadratic risk

$$\begin{aligned} \mathbb{E}(T_n - \mu)^2 &= \mathbb{E} \left[ (\bar{X}_n - \mu)^2 1_{|\bar{X}_n| > n^{-1/4}} \right] + \mu^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right) \\ &= \mathbb{E} \left[ (\bar{X}_n - \mu)^2 \right] - \mathbb{E} \left[ (\bar{X}_n - \mu)^2 1_{|\bar{X}_n| \leq n^{-1/4}} \right] + \mu^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right), \end{aligned}$$

we find that the second and third term decrease much faster than the first one. Hence the conclusion.  $\square$

Solution 1.5: we lower bound the MSE

$$\mathbb{E}(T_n - \mu_n)^2 = \mathbb{E} \left[ (\bar{X}_n - \mu_n)^2 1_{|\bar{X}_n| > n^{-1/4}} \right] + \mu_n^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right) \geq \mu_n^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right)$$

as the first term decreases at least as fast as  $O(1/n)$ . Since we have by the concentration inequality

$$\mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right) \geq \mathbb{P} \left( |\bar{X}_n - \mu_n| \leq \frac{1}{2}n^{-1/4} \right) \geq 1 - 2 \exp \left( -\frac{\sqrt{n}}{8b^2} \right),$$

it follows immediately that

$$\lim_{n \rightarrow \infty} \mu_n^2 \mathbb{P} \left( |\bar{X}_n| \leq n^{-1/4} \right) = \frac{1}{4\sqrt{n}}$$

hence the result.  $\square$

Solution 2.2: Let  $\eta$  denote the LRT's threshold and we want it to satisfy

$$\mathbb{P} \left( (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} Y > \eta | X = 0 \right) = \alpha.$$

Note that under  $X = 0$ , the test statistic is

$$(\mathbf{b} - \mathbf{a})^T \Sigma^{-1} Y = (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{a} + N) \sim \mathcal{N} \left( (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} \mathbf{a}, (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{b} - \mathbf{a}) \right).$$

Therefore, let  $\gamma := \sqrt{(\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{b} - \mathbf{a})}$ . The rejection region is

$$\left\{ \mathbf{y} \in \mathbb{R}^d, (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{y} - \mathbf{a}) \geq \gamma F^{-1}(1 - \alpha) \right\}$$

where  $F^{-1}(\cdot)$  denotes the quantile of the standard normal distribution.  $\square$

Solution 2.3: again, let us define the rejection region

$$R = \{ \mathbf{y} \in \mathbb{R}^d, (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} Y > \eta \}.$$

The type I error associated with the threshold  $\eta$  is

$$\mathbb{P} \left( (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (Y - \mathbf{a}) > \eta - (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} \mathbf{a} \mid X = 0 \right)$$

and its type II error is

$$\mathbb{P} \left( (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (Y - \mathbf{b}) \leq \eta - (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} \mathbf{b} \mid X = 1 \right)$$

It follows from the requirement that the two errors are equal

$$\eta - (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} \mathbf{b} = -\eta + (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} \mathbf{a}$$

that is

$$\eta = \frac{1}{2} (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{b} + \mathbf{a}). \quad \square$$

Solution 2.4: it suffices to look at the type I error, which can be written as

$$\mathbb{P} \left( (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} N > \frac{1}{2} (\mathbf{b} - \mathbf{a})^T \Sigma^{-1} (\mathbf{b} - \mathbf{a}) \right) = 1 - F \left( \frac{\gamma}{2} \right).$$

Hence, the stronger the SNR, the smaller the error.  $\square$